

8.1

$$Y = XB + U, \quad X \text{ } m \times p, \quad \text{rank } X = p, \quad B \text{ } p \times d$$

$$EU = 0$$

$$\hat{B} = (X^T X)^{-1} X^T Y \quad \text{cf. page 400 line 1}$$

$$(Y - X\hat{B})^T (Y - X\hat{B}) \leq (Y - XB)^T (Y - XB) \quad \text{cf. formula (8.10)}$$

$$\Rightarrow \text{tr}((Y - X\hat{B})^T (Y - X\hat{B})) \leq \text{tr}((Y - XB)^T (Y - XB)) \quad \text{cf. A 7.9 (6)}$$

$$\text{tr}((Y - XB)^T (Y - XB)) = \text{tr}((Y - X\hat{B})^T (Y - X\hat{B}))$$

$$\Leftrightarrow \text{tr}((Y - XB)^T (Y - XB) - (Y - X\hat{B})^T (Y - X\hat{B})) = 0$$

$$\Leftrightarrow \text{tr}((XB - X\hat{B})^T (XB - X\hat{B})) = 0 \quad \text{cf. formula (8.10) rearranged}^*$$

$$\Leftrightarrow \text{tr}((B - \hat{B})^T X^T X (B - \hat{B})) = 0, \quad X^T X > 0 \quad \text{cf. A 5.7}$$

$$\Leftrightarrow B - \hat{B} = 0$$

$$\Leftrightarrow B = \hat{B}$$

i. e. equality only for $B = \hat{B}$

$$* \quad (Y - \Theta)^T (Y - \Theta) = (Y - \hat{\Theta})^T (Y - \hat{\Theta}) + (\hat{\Theta} - \Theta)^T (\hat{\Theta} - \Theta)$$

$$\Leftrightarrow (Y - \Theta)^T (Y - \Theta) - (Y - \hat{\Theta})^T (Y - \hat{\Theta}) = (\Theta - \hat{\Theta})^T (\Theta - \hat{\Theta})$$

8.2

 $y_i \sim N_d(\theta_i, \Sigma)$, $i=1,2,3$, independent

$$\theta_1 = \beta_1 \quad \theta_2 = \beta_1 - \beta_2 \quad \theta_3 = \beta_1 + \beta_2$$

$$Y = X\beta + u = \begin{bmatrix} 1 & 0 \\ 1 & -1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} + u$$

$$\hat{\beta} = (X^T X)^{-1} X^T Y$$

$$X^T X = \begin{bmatrix} 1 & 1 & 1 \\ 0 & -1 & 1 \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 1 & -1 \\ 1 & 1 \end{bmatrix} = \begin{bmatrix} 3 & 0 \\ 0 & 2 \end{bmatrix}, \quad (X^T X)^{-1} = \frac{1}{6} \begin{bmatrix} 2 & 0 \\ 0 & 3 \end{bmatrix}$$

$$(X^T X)^{-1} X^T = \frac{1}{6} \begin{bmatrix} 2 & 0 \\ 0 & 3 \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 \\ 0 & -1 & 1 \end{bmatrix} = \begin{bmatrix} \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ 0 & -\frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

$$\hat{\beta} = \begin{bmatrix} \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ 0 & -\frac{1}{2} & \frac{1}{2} \end{bmatrix} Y, \quad \hat{\beta}^T = Y^T \begin{bmatrix} \frac{1}{3} & 0 \\ \frac{1}{3} & -\frac{1}{2} \\ \frac{1}{3} & \frac{1}{2} \end{bmatrix}$$

$$\hat{\beta}_1 = \frac{1}{3} (y_1 + y_2 + y_3)$$

$$\hat{\beta}_2 = \frac{1}{2} (y_3 - y_2)$$

$$\hat{\Theta} = X\hat{\beta} = \begin{bmatrix} 1 & 0 \\ 1 & -1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ 0 & -\frac{1}{2} & \frac{1}{2} \end{bmatrix} Y = \begin{bmatrix} \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ \frac{1}{3} & \frac{2}{6} & -\frac{1}{2} \\ \frac{1}{3} & -\frac{1}{6} & \frac{5}{6} \end{bmatrix} Y$$

$$\hat{\Theta}^T = Y^T \begin{bmatrix} \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ \frac{1}{3} & \frac{2}{6} & -\frac{1}{2} \\ \frac{1}{3} & -\frac{1}{6} & \frac{5}{6} \end{bmatrix}$$

$$\hat{\theta}_1 = \frac{1}{3} (y_1 + y_2 + y_3) \quad (= \hat{\beta}_1)$$

$$\hat{\theta}_2 = \frac{1}{3} y_1 + \frac{2}{6} y_2 - \frac{1}{2} y_3 \quad (= \hat{\beta}_1 - \hat{\beta}_2)$$

$$\hat{\theta}_3 = \frac{1}{3} y_1 - \frac{1}{6} y_2 + \frac{5}{6} y_3 \quad (= \hat{\beta}_1 + \hat{\beta}_2)$$

8.3

 $\underline{y}_i \sim N_d(\underline{\mu}_i, \Sigma)$, $i=1, \dots, n$, independent

$$\underline{\mu}_i = B^T \underline{x}_i, \quad B \quad n \times d$$

$$X = [\underline{x}_1 \quad \underline{x}_2 \quad \dots \quad \underline{x}_n]^T \quad n \times n, \quad \text{rank } X = n$$

 $\underline{y}_i = B^T \underline{x}_i + \underline{u}_i$, $\underline{u}_i \sim N_d(\underline{0}, \Sigma)$, $i=1, \dots, n$, independent

 $Y = XB + U$, i.e. the usual model

Consider $\hat{y}_0 = \hat{B}^T \underline{x}_0$, $\hat{B} = (X^T X)^{-1} X^T Y$

$$\hat{y}_{0j} = \hat{\beta}^{(j)T} \underline{x}_0 = \underline{x}_0^T \hat{\beta}^{(j)}$$

$$\begin{aligned} \text{Cov}(\hat{y}_{0j}, \hat{y}_{0k}) &= \text{Cov}(\underline{x}_0^T \hat{\beta}^{(j)}, \underline{x}_0^T \hat{\beta}^{(k)}) = \underline{x}_0^T \text{Cov}(\hat{\beta}^{(j)}, \hat{\beta}^{(k)}) \underline{x}_0 \\ &= \underline{x}_0^T c_{jk} (X^T X)^{-1} \underline{x}_0 = \underline{x}_0^T (X^T X)^{-1} \underline{x}_0 \sigma_{jk} \\ &\quad \text{cf. formula (8.22)} \end{aligned}$$

Hence $\text{Var } \hat{y}_0 = \underline{x}_0^T (X^T X)^{-1} \underline{x}_0 \Sigma$

As Σ is estimated by $S (= \frac{E}{n-p})$ we see that $\text{Var } \hat{y}_0$ is estimated by $\underline{x}_0^T (X^T X)^{-1} \underline{x}_0 S$

Consider $\underline{a}^T \hat{y}_0 = \underline{a}^T \hat{B}^T \underline{x}_0$, lin. comb. of normal dist. var.

$$E[\underline{a}^T \hat{y}_0] = \underline{a}^T E \hat{B}^T \underline{x}_0 = \underline{a}^T B^T \underline{x}_0 = \underline{a}^T \underline{\mu}_0$$

$$\text{Var}(\underline{a}^T \hat{y}_0) = \underline{a}^T \underline{x}_0^T (X^T X)^{-1} \underline{x}_0 \Sigma \underline{a} = \underline{x}_0^T (X^T X)^{-1} \underline{x}_0 \underline{a}^T \Sigma \underline{a}$$

$$\underline{a}^T \hat{y}_0 \sim N(\underline{a}^T \underline{\mu}_0, \underline{x}_0^T (X^T X)^{-1} \underline{x}_0 \underline{a}^T \Sigma \underline{a})$$

$$\frac{\underline{a}^T \hat{y}_0 - \underline{a}^T \underline{\mu}_0}{\sqrt{\underline{x}_0^T (X^T X)^{-1} \underline{x}_0 \underline{a}^T S \underline{a}}} \sim t(n-p)$$

Hence the confidence interval for $\underline{a}^T \underline{\mu}_0$ with confidence level $1-\alpha$

$$\underline{a}^T \underline{\mu}_0 = \underline{a}^T \hat{y}_0 \pm t_{1-\frac{\alpha}{2}}(n-p) \sqrt{\underline{x}_0^T (X^T X)^{-1} \underline{x}_0 \underline{a}^T S \underline{a}}$$

cont.

8.3 continued

Let y_0 now indicate a new observation at x_0 independent of the earlier observations y_1, \dots, y_n

$$y_0 = \beta^T x_0 + u_0, \quad u_0 \sim N_d(0, \Sigma)$$

Consider $\underline{a}^T y_0 - \underline{a}^T \hat{y}_0 = \underline{a}^T y_0 - \underline{a}^T \hat{\beta}^T x_0$, lin. comb. of normal dist. var.

$$E[\underline{a}^T y_0 - \underline{a}^T \hat{y}_0] = \underline{a}^T \beta^T x_0 - \underline{a}^T \beta^T x_0 = 0$$

$$\begin{aligned} \text{Var}(\underline{a}^T y_0 - \underline{a}^T \hat{y}_0) &= \underline{a}^T \Sigma \underline{a} + x_0^T (X^T X)^{-1} x_0 \underline{a}^T \Sigma \underline{a} \\ &= (1 + x_0^T (X^T X)^{-1} x_0) \underline{a}^T \Sigma \underline{a} \end{aligned}$$

$$\underline{a}^T y_0 - \underline{a}^T \hat{y}_0 \sim N(0, (1 + x_0^T (X^T X)^{-1} x_0) \underline{a}^T \Sigma \underline{a})$$

$$\frac{\underline{a}^T y_0 - \underline{a}^T \hat{y}_0 - 0}{\sqrt{(1 + x_0^T (X^T X)^{-1} x_0) \underline{a}^T \Sigma \underline{a}}} \sim t(n-p)$$

Hence the prediction interval for $\underline{a}^T y_0$ with confidence level $1 - \alpha$

$$\underline{a}^T y_0 = \underline{a}^T \hat{y}_0 \pm t_{1-\frac{\alpha}{2}}(n-p) \sqrt{(1 + x_0^T (X^T X)^{-1} x_0) \underline{a}^T \Sigma \underline{a}}$$

8.4

$y_i \sim N_d(\underline{\mu}, \Sigma)$, $i=1, \dots, n$, independent

$$Y = \underline{1}_d \underline{\mu}^T + U, \quad u_i \sim N_d(0, \Sigma), \quad i=1, \dots, n, \text{ indep.}$$

$$\hat{\underline{\mu}}^T = (\underline{1}_d^T \underline{1}_d)^{-1} \underline{1}_d^T Y = \frac{1}{d} \sum_{i=1}^d y_i^T = \bar{y}^T$$

$$\hat{\underline{\mu}} = \bar{y}$$