

8.9

The usual linear normal model: $Y = XB + U$

In exercise 8.3 we found the confidence interval for

$$\underline{\underline{a}}^T \underline{\underline{\mu}}_0 = \underline{\underline{a}}^T B^T \underline{\underline{x}}_0 \text{ as}$$

$$\underline{\underline{a}}^T \underline{\underline{\mu}}_0 = \underline{\underline{a}}^T \hat{\underline{\underline{y}}}_0 \pm t_{1-\frac{\alpha}{2}}(n-p) \sqrt{\underline{\underline{x}}_0^T (X^T X)^{-1} \underline{\underline{x}}_0 \underline{\underline{a}}^T S \underline{\underline{a}}}$$

With $\underline{\underline{a}} := \underline{\underline{b}}^T$ and $\underline{\underline{x}}_0 := \underline{\underline{a}}$ we find $\underline{\underline{\mu}}_0 = B^T \underline{\underline{a}}$ and $\hat{\underline{\underline{y}}}_0 = \hat{B}^T \underline{\underline{a}}$:

$$\underline{\underline{b}}^T B^T \underline{\underline{a}} = \underline{\underline{b}}^T \hat{B}^T \underline{\underline{a}} \pm t_{1-\frac{\alpha}{2}}(n-p) \sqrt{\underline{\underline{a}}^T (X^T X)^{-1} \underline{\underline{a}} \underline{\underline{b}}^T S \underline{\underline{b}}}$$

Note that $\underline{\underline{b}}^T B^T \underline{\underline{a}} = \underline{\underline{a}}^T B \underline{\underline{b}}$ and $\underline{\underline{b}}^T \hat{B}^T \underline{\underline{a}} = \underline{\underline{a}}^T \hat{B} \underline{\underline{b}}$. The confidence interval for $\underline{\underline{a}}^T B \underline{\underline{b}}$ with confidence level $1-\alpha$ then becomes

$$\underline{\underline{a}}^T B \underline{\underline{b}} = \underline{\underline{a}}^T \hat{B} \underline{\underline{b}} \pm t_{1-\frac{\alpha}{2}}(n-p) \sqrt{\underline{\underline{a}}^T (X^T X)^{-1} \underline{\underline{a}} \underline{\underline{b}}^T S \underline{\underline{b}}}$$

Considering in pre specified confidence intervals of the type above (i.e. for $\underline{\underline{a}}^T B \underline{\underline{b}}$) with simultaneous confidence level at least $1-\alpha$, we find according to Bonferroni's method that

$$\underline{\underline{a}}^T B \underline{\underline{b}} = \underline{\underline{a}}^T \hat{B} \underline{\underline{b}} \pm t_{1-\frac{\alpha}{2m}}(n-p) \sqrt{\underline{\underline{a}}^T (X^T X)^{-1} \underline{\underline{a}} \underline{\underline{b}}^T S \underline{\underline{b}}}$$

The confidence interval found in formula (8.63) was

$$\underline{\underline{a}}^T B \underline{\underline{b}} = \underline{\underline{a}}^T \hat{B} \underline{\underline{b}} \pm \sqrt{(n-p) \varphi_{1-\alpha} \underline{\underline{a}}^T (X^T X)^{-1} \underline{\underline{a}} \underline{\underline{b}}^T S \underline{\underline{b}}}$$

8.10

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$$Y = XB + U, \quad \underline{u}_i \sim N_d(\underline{0}, \Sigma), \quad i=1, \dots, n, \quad \text{independent}$$

$$X = [X_1 \ X_2], \quad X \ n \times n, \quad \text{rank } X = p$$

$$X_1 \ n \times r_1, \quad X_2 \ n \times r_2, \quad r_1 + r_2 = n$$

$$B = \begin{bmatrix} B_1 \\ B_2 \end{bmatrix}, \quad B \ p \times d$$

$$B_1 \ r_1 \times d, \quad B_2 \ r_2 \times d$$

$$XB = X_1 B_1 + X_2 B_2$$

$$\hat{B} = (X^T X)^{-1} X^T Y$$

$$E = Y^T (I_n - X(X^T X)^{-1} X^T) Y$$

$$(i) H_0: B_2 = 0 \Leftrightarrow B_H = \begin{bmatrix} B_1 \\ 0 \end{bmatrix}$$

$$\hat{B}_H = \begin{bmatrix} \hat{B}_1 \\ 0 \end{bmatrix} = \begin{bmatrix} (X_1^T X_1)^{-1} X_1^T Y \\ 0 \end{bmatrix}$$

$$X \hat{B}_H = [X_1 \ X_2] \begin{bmatrix} (X_1^T X_1)^{-1} X_1^T Y \\ 0 \end{bmatrix} = X_1 (X_1^T X_1)^{-1} X_1^T Y$$

$$E_H = Y^T (I_n - X_1 (X_1^T X_1)^{-1} X_1^T) Y$$

$$\Lambda = \frac{\det E}{\det(E+H)} = \frac{\det E}{\det E_H} \sim U(d, r_2, n-d) \text{ under } H_0$$

$$H = E_H - E = Y^T (X(X^T X)^{-1} X^T - X_1 (X_1^T X_1)^{-1} X_1^T) Y$$

$$(ii) H_0: B_2 = 0 \Leftrightarrow \begin{bmatrix} 0 & I_{r_2} \end{bmatrix} \begin{bmatrix} B_1 \\ B_2 \end{bmatrix} = 0 \Leftrightarrow AB = 0$$

$$\omega = \mathcal{N}(A_1), \quad A_1 = A(X^T X)^{-1} X^T = \begin{bmatrix} 0 & I_{r_2} \end{bmatrix} (X^T X)^{-1} X^T$$

$$= \mathcal{R}(A_1^T)^\perp = \mathcal{R}([X_1 \ 0]) \approx A_1^T [X_1 \ 0] = X(X^T X)^{-1} \begin{bmatrix} 0 \\ I_{r_2} \end{bmatrix} [X_1 \ 0] = 0$$

$$P_\omega - P_\omega = X(X^T X)^{-1} X^T - X_1 (X_1^T X_1)^{-1} X_1^T$$

$$H = Y^T (P_\omega - P_\omega) Y = Y^T (X(X^T X)^{-1} X^T - X_1 (X_1^T X_1)^{-1} X_1^T) Y$$

thus $E+H$ as well as Λ can be found without first specifying \hat{B}_H .

Determination of H from "corollary" 1 to theorem 8.5 and determination of H by use of $P_{\omega+n, \omega}$ both involve calculation of $(X^T X)^{-1} A^T (A (X^T X)^{-1} A^T)^{-1} A (X^T X)^{-1}$.

(iii)

$$\begin{aligned} H &= (A \hat{B})^T (A (X^T X)^{-1} A^T)^{-1} A \hat{B} \\ &= ([0 \ I_{n_2}] (X^T X)^{-1} X^T Y)^T ([0 \ I_{n_2}] (X^T X)^{-1} [0 \ I_{n_2}]^T)^{-1} [0 \ I_{n_2}] (X^T X)^{-1} X^T Y \\ &= Y^T X G X^T Y, \quad G = (X^T X)^{-1} \begin{bmatrix} 0 \\ I_{n_2} \end{bmatrix} ([0 \ I_{n_2}] (X^T X)^{-1} \begin{bmatrix} 0 \\ I_{n_2} \end{bmatrix})^{-1} [0 \ I_{n_2}] (X^T X)^{-1} \end{aligned}$$

Using A3.1: $\begin{bmatrix} A & B \\ B^T & D \end{bmatrix}^{-1} = \begin{bmatrix} A^{-1} + FE^{-1}F^T & -FE^{-1} \\ -E^{-1}F^T & E^{-1} \end{bmatrix}$, $E = D - B^T A^{-1} B$
 $F = A^{-1} B$

Note that

$$\begin{aligned} &\begin{bmatrix} A & B \\ B^T & D \end{bmatrix}^{-1} \begin{bmatrix} 0 \\ I_{n_2} \end{bmatrix} ([0 \ I_{n_2}] \begin{bmatrix} A & B \\ B^T & D \end{bmatrix}^{-1} \begin{bmatrix} 0 \\ I_{n_2} \end{bmatrix})^{-1} [0 \ I_{n_2}] \begin{bmatrix} A & B \\ B^T & D \end{bmatrix}^{-1} \\ &= \begin{bmatrix} -FE^{-1} \\ E^{-1} \end{bmatrix} (E^{-1})^{-1} [-E^{-1}F^T \ E^{-1}] = \begin{bmatrix} -F \\ I_{n_2} \end{bmatrix} [-E^{-1}F^T \ E^{-1}] \\ &= \begin{bmatrix} FE^{-1}F^T & -FE^{-1} \\ -E^{-1}F^T & E^{-1} \end{bmatrix} = \begin{bmatrix} A^{-1} + FE^{-1}F^T & -FE^{-1} \\ -E^{-1}F^T & E^{-1} \end{bmatrix} - \begin{bmatrix} A^{-1} & 0 \\ 0 & 0 \end{bmatrix} \\ &= \begin{bmatrix} A & B \\ B^T & D \end{bmatrix}^{-1} - \begin{bmatrix} A^{-1} & 0 \\ 0 & 0 \end{bmatrix} \end{aligned}$$

Hence $G = (X^T X)^{-1} - \begin{bmatrix} (X_1^T X_1)^{-1} & 0 \\ 0 & 0 \end{bmatrix}$

$$H = Y^T X G X^T Y$$

$$= Y^T (X (X^T X)^{-1} X^T - [X_1 \ X_2] \begin{bmatrix} (X_1^T X_1)^{-1} & 0 \\ 0 & 0 \end{bmatrix} \begin{bmatrix} X_1^T \\ X_2^T \end{bmatrix}) Y$$

$$= Y^T (X (X^T X)^{-1} X^T - X_1 (X_1^T X_1)^{-1} X_1^T) Y$$

(iv)

$$P_{\omega+n, \omega} = X (X^T X)^{-1} A^T (A (X^T X)^{-1} A^T)^{-1} A (X^T X)^{-1} X^T = X G X^T$$

$$= X ((X^T X)^{-1} - \begin{bmatrix} (X_1^T X_1)^{-1} & 0 \\ 0 & 0 \end{bmatrix}) X^T = X (X^T X)^{-1} X^T - X_1 (X_1^T X_1)^{-1} X_1^T$$

$$H = Y^T P_{\omega+n, \omega} Y = Y^T (X (X^T X)^{-1} X^T - X_1 (X_1^T X_1)^{-1} X_1^T) Y$$

$$H_{01}: A_1 B C_1^T = 0 \Leftrightarrow C_1 (\underline{\mu}_{\tilde{k}-1} - \underline{\mu}_{\tilde{k}}) = \underline{0}, \quad k=2, \dots, K$$

$$H_{03}: \mathbf{1}_K^T B C_1^T = \underline{0}^T \Leftrightarrow C_1 \sum_{k=1}^K \underline{\mu}_{\tilde{k}} = \underline{0}$$

$$B = \begin{bmatrix} \underline{\mu}_{\tilde{1}}^T \\ \underline{\mu}_{\tilde{2}}^T \\ \vdots \\ \underline{\mu}_{\tilde{K}}^T \end{bmatrix}, \quad \left. \begin{matrix} A_1 \\ C_1 \end{matrix} \right\} = \begin{bmatrix} 1 & -1 & 0 & \dots & 0 \\ 0 & 1 & -1 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & -1 \end{bmatrix} \quad \begin{matrix} A_1: (K-1) \times K \\ C_1: (d-1) \times d \\ \text{rank } A_1 = K-1 \\ \text{rank } C_1 = d-1 \end{matrix}$$

$$\underline{x}_{\tilde{i}}^k \sim N_d(\underline{\mu}_{\tilde{k}}, \Sigma), \quad k=1, \dots, K, \quad i=1, \dots, m_k, \quad \text{independent}$$

$$\bar{\underline{x}}^k = \frac{1}{m_k} \sum_{i=1}^{m_k} \underline{x}_{\tilde{i}}^k, \quad k=1, \dots, K$$

$$S_k = \frac{1}{m_k-1} Q_k, \quad Q_k = \sum_{i=1}^{m_k} (\underline{x}_{\tilde{i}}^k - \bar{\underline{x}}^k)(\underline{x}_{\tilde{i}}^k - \bar{\underline{x}}^k)^T, \quad k=1, \dots, K, \quad \text{indep.}$$

$$\bar{\underline{x}}^k \sim N_d\left(\underline{\mu}_{\tilde{k}}, \frac{\Sigma}{m_k}\right), \quad k=1, \dots, K$$

$$(m_k-1) S_k \sim W_d(m_k-1, \Sigma), \quad k=1, \dots, K$$

} independent

$$\bar{\underline{x}} = \frac{1}{m} \sum_{k=1}^K m_k \bar{\underline{x}}^k, \quad m = \sum_{k=1}^K m_k$$

$$S_n = \frac{1}{m-K} \sum_{k=1}^K (m_k-1) S_k$$

$$\bar{\underline{x}} \sim N_d\left(\frac{1}{m} \sum_{k=1}^K m_k \underline{\mu}_{\tilde{k}}, \frac{\Sigma}{m}\right)$$

$$(m-K) S_n \sim W_d(m-K, \Sigma)$$

} independent

$$C_1 \bar{\underline{x}} \sim N_{d-1}\left(\frac{1}{m} \sum_{k=1}^K m_k C_1 \underline{\mu}_{\tilde{k}}, \frac{1}{m} C_1 \Sigma C_1^T\right)$$

$$(m-K) S_n \sim W_{d-1}(m-K, C_1 \Sigma C_1^T)$$

} independent

Note that $C_1 \underline{\mu}_{\tilde{k}-1} = C_1 \underline{\mu}_{\tilde{k}}, k=2, \dots, K \Rightarrow C_1 \underline{\mu}_{\tilde{k}} = \frac{1}{K} \sum_{k=1}^K C_1 \underline{\mu}_{\tilde{k}}, k=1, \dots, K$

$$C_1 \bar{\underline{x}} \sim N_{d-1}\left(\frac{1}{K} \sum_{k=1}^K C_1 \underline{\mu}_{\tilde{k}}, \frac{1}{m} C_1 \Sigma C_1^T\right) \quad \text{when } H_{01} \text{ is true}$$

$$C_1 \bar{\underline{x}} \sim N_{d-1}(\underline{0}, \frac{1}{m} C_1 \Sigma C_1^T) \quad \text{when } H_{01} \cap H_{03} \text{ is true}$$

$$T^2 = m(m-K) (C_1 \bar{\underline{x}})^T ((m-K) C_1 S_n C_1^T)^{-1} C_1 \bar{\underline{x}}$$

$$= m \bar{\underline{x}}^T C_1^T (C_1 S_n C_1^T)^{-1} C_1 \bar{\underline{x}}$$

$$\sim T^2(d-1, m-K) \quad \text{when } H_{01} \cap H_{03} \text{ is true, cf. corollary to th. 2.2}$$

$$\begin{aligned}
 [y^{(1)} \ y^{(2)}] &= [X_1 \beta^{(1)} \ X_2 \beta^{(2)}] + [\underline{u}^{(1)} \ \underline{u}^{(2)}] \\
 &= [X_1 \ X_2] \begin{bmatrix} \beta^{(1)} \\ \beta^{(2)} \end{bmatrix} + [\underline{u}^{(1)} \ \underline{u}^{(2)}] \\
 &= X \beta + U
 \end{aligned}$$

$$X = \begin{bmatrix} 1 & x_1 & 1 & x_1 & x_1^2 \\ 1 & x_2 & 1 & x_2 & x_2^2 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & x_n & 1 & x_n & x_n^2 \end{bmatrix}, \quad \beta^T = [\beta_{10} \ \beta_{11} \ \beta_{20} \ \beta_{21} \ \beta_{22}]$$

$$n \geq 3$$

$$H_0: A \beta = \underline{0}, \quad A \text{ } q \times 5, \quad q \leq 5$$

Presuming that at least three of the x_i 's are different, then $\text{rank } X_2 = 3$, hence $\text{rank } X = 3$. The rows of X therefore span \mathbb{R}^3 . Thus if only $a_{j1} = a_{j3} \wedge a_{j2} = a_{j4}$, $j=1, \dots, q$, then $(\underline{a}_j^T)^T \beta$, $j=1, \dots, q$, are estimable.

$$a \quad H_0: \beta_{11} = \beta_{21} = \beta_{22} = 0$$

$$\Leftrightarrow H_{01}: \beta_{11} = 0 \quad \wedge \quad H_{02}: \beta_{21} = \beta_{22} = 0$$

$$\Leftrightarrow H_{01}: [0 \ 1] \beta^{(1)} = 0 \quad \wedge \quad H_{02}: \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} \beta^{(2)} = \underline{0}$$

$$\Leftrightarrow H_0: \begin{bmatrix} 0 & 1 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{bmatrix} \beta = \underline{0}$$

$$(\underline{a}_j^T)^T \beta, \quad j=1, 2, \text{ are estimable}$$

$$b \quad H_0: \beta_{22} = 0$$

$$\Leftrightarrow H_{01}: \text{no restrictions} \quad \wedge \quad H_{02}: \beta_{22} = 0$$

$$\Leftrightarrow H_{01}: [0 \ 0] \beta_{\sim}^{(1)} = 0 \wedge H_{02}: [0 \ 0 \ 1] \beta_{\sim}^{(2)} = 0$$

$$\Leftrightarrow H_0: [0 \ 0 \ 0 \ 0 \ 1] \beta_{\sim} = 0$$

$A\beta_{\sim}$ is estimable

c
$$H_0: \beta_{11} = \beta_{21} = 0$$

$$\Leftrightarrow H_{01}: \beta_{11} = 0 \wedge H_{02}: \beta_{21} = 0$$

$$\Leftrightarrow H_{01}: [0 \ 1] \beta_{\sim}^{(1)} = 0 \wedge H_{02}: [0 \ 1 \ 0] \beta_{\sim}^{(2)} = 0$$

$$\Leftrightarrow H_0: [0 \ 1 \ 0 \ 1 \ 0] \beta_{\sim} = 0$$

$A\beta_{\sim}$ is estimable

d
$$H_0: \beta_{11} = 0$$

$$\Leftrightarrow H_{01}: \beta_{11} = 0 \wedge H_{02}: \text{no restrictions}$$

$$\Leftrightarrow H_{01}: [0 \ 1] \beta_{\sim}^{(1)} = 0 \wedge H_{02}: [0 \ 0 \ 0] \beta_{\sim}^{(2)} = 0$$

$$\Leftrightarrow H_0: [0 \ 1 \ 0 \ 0 \ 0] \beta_{\sim} = 0$$

$A\beta_{\sim}$ is not estimable, as $a_{12} \neq a_{14}$

e
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$$\Leftrightarrow H_{01}: [0 \ 1] \beta_{\sim}^{(1)} = 0 \wedge H_{02}: [0 \ 0 \ 1] \beta_{\sim}^{(2)} = 0$$

$$\Leftrightarrow H_0: [0 \ 1 \ 0 \ 0 \ 1] \beta_{\sim} = 0$$

$A\beta_{\sim}$ is not estimable, as $a_{12} \neq a_{14}$