

$$W \sim W_d(m, \Sigma), \quad m \geq d$$

$$\text{Let } W = \begin{bmatrix} W_{11} & W_{12} \\ W_{21} & W_{22} \end{bmatrix} \quad \begin{array}{l} W_{11} \quad d_1 \times d_1 \\ W_{22} \quad d_2 \times d_2 \end{array}$$

$$\text{and } W_{22.1} = W_{22} - W_{21} W_{11}^{-1} W_{12}$$

Lemma 2.10

$$W_{22.1} \sim W_{d_2}(m - d_1, \Sigma_{22.1}) \quad \text{indep. of } (W_{11}, W_{12})$$

Proof =

$$\text{Assume } W = \sum_i \tilde{x}_i \tilde{x}_i^T, \quad \tilde{x}_i \sim N_d(\underline{0}, \Sigma), \quad i=1, \dots, m$$

independent

$$\text{Let } \tilde{x}_i = \begin{bmatrix} u_i \\ v_i \end{bmatrix}, \quad u_i \text{ } d_1\text{-dim.}, \quad v_i \text{ } d_2\text{-dim.}$$

$$W = \begin{bmatrix} \sum_i u_i u_i^T & \sum_i u_i v_i^T \\ \sum_i v_i u_i^T & \sum_i v_i v_i^T \end{bmatrix} = \begin{bmatrix} U^T U & U^T V \\ V^T U & V^T V \end{bmatrix}$$

$$\begin{aligned} W_{22.1} &= V^T V - V^T U (U^T U)^{-1} U^T V \\ &= V^T (I_m - U (U^T U)^{-1} U^T) V \\ &= V^T (I_m - P) V \end{aligned}$$

Note that $P|U$ is a projection matrix

$$\text{rank } P = \text{rank } U = d_1, \text{ a.s.}$$

$$\text{Let } \tilde{v}_{i.1} = \tilde{v}_i - \sum_{21} \Sigma_{11}^{-1} u_i$$

$$V_{2.1} = \begin{bmatrix} \underline{v}_{1.1}^T \\ \vdots \\ \underline{v}_{m-1.1}^T \end{bmatrix} = V - U \Sigma_{11}^{-1} \Sigma_{12}$$

Note that $(I_m - P) V_{2.1} = (I_m - P) (V - U \Sigma_{11}^{-1} \Sigma_{12})$
 $= (I_m - P) V - (U - U) \Sigma_{11}^{-1} \Sigma_{12}$
 $= (I_m - P) V$

$$\Rightarrow W_{22.1} = V_{2.1}^T (I_m - P) V_{2.1}$$

$$\underline{v}_{i.1} | \underline{u}_i \sim N_{d_2}(\underline{0} + \Sigma_{21} \Sigma_{11}^{-1} (\underline{u}_i - \underline{0}), \Sigma_{22.1}) \quad \text{cf. th. 2.1 (viii)}$$

$$= N_{d_2}(\Sigma_{21} \Sigma_{11}^{-1} \underline{u}_i, \Sigma_{22.1}), \quad i = 1, \dots, m$$

independent

$$\Rightarrow \underline{v}_{i-1.1} | \underline{u}_i \sim N_{d_2}(\underline{0}, \Sigma_{22.1}), \quad i = 1, \dots, m$$

independent

$$\Rightarrow V_{2.1}^T (I_m - P) V_{2.1} | U \sim W_{d_2}(m - d_1, \Sigma_{22.1})$$

cf. th 2.4 corollary 1

The distribution does not depend on U

$$\Rightarrow V_{2.1}^T (I_m - P) V_{2.1} \sim W_{d_2}(m - d_1, \Sigma_{22.1})$$

Note that $P(I_m - P)^T = P(I_m - P) = P - P = 0$

$$\Rightarrow \left\{ \begin{array}{l} P V_{2.1} I_{d_2} | U \\ (I_m - P) V_{2.1} I_{d_2} | U \end{array} \right\} \text{ independent (cf. ex. 2.18)}$$

$$\Rightarrow P V_{2.1} | U \text{ and } W_{22.1} | U \text{ independent}$$

$$\Rightarrow (P V_{2.1}, U) \text{ and } W_{22.1} \text{ independent,}$$

as $x|u$ and $y|u$ indep. \wedge y and u indep.

$$\Rightarrow f(x, y | u) = f(x|u) f(y|u) = f(x|u) f(y)$$

$$\Rightarrow f(x, y, u) = f(x, u) f(y)$$

$\Rightarrow (x, u)$ and y independent

Note that $W_{11} = u^T u$

$$\begin{aligned} W_{12} &= u^T V = u^T P V = u^T P (V_{2.1} + u \Sigma_{11}^{-1} \Sigma_{12}) \\ &= u^T P V_{2.1} + u^T u \Sigma_{11}^{-1} \Sigma_{12} \end{aligned}$$

that is W_{11} and W_{12} are both functions of $(P V_{2.1}, u)$

thus $W_{22.1}$ and (W_{11}, W_{12}) are independent

Note that $W_{21} W_{11}^{-1} W_{12}$ is independent of $W_{22.1}$,
as $W_{21} W_{11}^{-1} W_{12}$ is a function of (W_{11}, W_{12}) .

Corollary:

$$\Sigma_{12} = 0 \Rightarrow W_{21} W_{11}^{-1} W_{12} \sim W_{d_2}(d_1, \Sigma_{22})$$

Proof:

$$\Sigma_{12} = 0 \Rightarrow \Sigma_{22.1} = \Sigma_{22} \wedge V_{2.1} = V$$

$$\begin{aligned} \text{Note that } W_{21} W_{11}^{-1} W_{12} &= W_{22} - W_{22.1} \\ &= V^T V - V^T (E_{\text{rank}} - P) V \\ &= V^T P V \\ &= V_{2.1}^T P V_{2.1} \end{aligned}$$

$$\Rightarrow W_{21} W_{11}^{-1} W_{12} | u \sim W_{d_2}(d_1, \Sigma_{22})$$

$$\Rightarrow W_{21} W_{11}^{-1} W_{12} \sim W_{d_2}(d_1, \Sigma_{22}) \text{ indep. of } u \text{ thus indep. of } W_{11}$$

$$\left. \begin{aligned} \underline{x} &\sim N_d(\underline{\mu}, \Sigma) \\ W &\sim W_d(m, \Sigma) \end{aligned} \right\} \text{independent}$$

Let $\underline{x} = \begin{bmatrix} \underline{x}^{(1)} \\ \underline{x}^{(2)} \end{bmatrix}$ and $\underline{\mu} = \begin{bmatrix} \underline{\mu}^{(1)} \\ \underline{\mu}^{(2)} \end{bmatrix}$ $\underline{x}^{(1)}, \underline{\mu}^{(1)}$ d_1 -dim.
 $\underline{x}^{(2)}, \underline{\mu}^{(2)}$ d_2 -dim.

$$\left. \begin{aligned} \Delta_d^2 &= \underline{\mu}^T \Sigma^{-1} \underline{\mu} \\ \Delta_{d_1}^2 &= \underline{\mu}^{(1)T} \Sigma_{11}^{-1} \underline{\mu}^{(1)} \end{aligned} \right\} \text{Mahalanobis' squared distances}$$

$$\Delta_d^2 - \Delta_{d_1}^2 = \underline{\mu}_{2..}^T \Sigma_{22..}^{-1} \underline{\mu}_{2..}, \quad \underline{\mu}_{2..} = \underline{\mu}^{(2)} - \Sigma_{21} \Sigma_{11}^{-1} \underline{\mu}^{(1)}$$

cf. ex. 2.20

Note that $\Delta_d^2 = \Delta_{d_1}^2 \Leftrightarrow \underline{\mu}_{2..} = \underline{0} \quad (\Sigma_{22..} > 0)$

Sample versions:

$$T_d^2 = m \underline{x}^T W^{-1} \underline{x}$$

$$T_{d_1}^2 = m \underline{x}^{(1)T} W_{11}^{-1} \underline{x}^{(1)}$$

$$T_d^2 - T_{d_1}^2 = m \underline{x}_{2..}^T W_{22..}^{-1} \underline{x}_{2..}, \quad \underline{x}_{2..} = \underline{x}^{(2)} - W_{21} W_{11}^{-1} \underline{x}^{(1)}$$

cf. ex. 2.20

Theorem 2.11

When $H_0: \underline{\mu}_{2..} = \underline{0}$ is true we have

$$\frac{T_d^2 - T_{d_1}^2}{m + T_{d_1}^2} \sim \frac{d_2}{m - d + 1} F(d_2, m - d + 1)$$

indep. of $T_{d_1}^2$ (only when H_0 is true)

Proof: Let $W = \begin{bmatrix} w_{11} & w_{12} \\ w_{21} & w_{22} \end{bmatrix} = \begin{bmatrix} u^T u & u^T v \\ v^T u & v^T v \end{bmatrix}$ $\begin{matrix} d_1 \\ d_2 \end{matrix}$

$$\begin{aligned} \underline{x}_{z,1}^{(2)} | \underline{x}_{z,1}^{(1)} &\sim N_{d_2} \left(\underline{\mu}_{z,1}^{(2)} + \Sigma_{z,1} \Sigma_{z,1}^{-1} (\underline{x}_{z,1}^{(1)} - \underline{\mu}_{z,1}^{(1)}), \Sigma_{z,2,1} \right) \\ &= N_{d_2} \left(\Sigma_{z,1} \Sigma_{z,1}^{-1} \underline{x}_{z,1}^{(1)}, \Sigma_{z,2,1} \right), H_0 \text{ true} \end{aligned}$$

$$\underline{x}_{z,1}^{(2)} = \underline{x}_{z,1}^{(2)} - w_{21} w_{11}^{-1} \underline{x}_{z,1}^{(1)}$$

Note that $w_{21} w_{11}^{-1} \underline{x}_{z,1}^{(1)} = v^T u w_{11}^{-1} \underline{x}_{z,1}^{(1)} = v^T \underline{a}_{z,1}$
 $= a_1 \underline{v}_1 + \dots + a_m \underline{v}_m$

$\underline{x}_{z,1}^{(2)} | (\underline{x}_{z,1}^{(1)}, U) \sim N_{d_2}$, as $\underline{x}_{z,1}^{(2)}$ is a lin. comb. of $m+1$ indep. d_2 -dim. normal distributed vectors ($\underline{x}_{z,1}^{(2)}, \underline{v}_1, \dots, \underline{v}_m$ independent)

Note that $E[w_{21} w_{11}^{-1} \underline{x}_{z,1}^{(1)} | (\underline{x}_{z,1}^{(1)}, U)]$
 $= E[v^T u w_{11}^{-1} \underline{x}_{z,1}^{(1)} | (\underline{x}_{z,1}^{(1)}, U)]$
 $= E[v^T | (\underline{x}_{z,1}^{(1)}, U)] u w_{11}^{-1} \underline{x}_{z,1}^{(1)}$
 $= \Sigma_{z,1} \Sigma_{z,1}^{-1} u^T u w_{11}^{-1} \underline{x}_{z,1}^{(1)}$, as $E[\underline{v}_i | \underline{u}_i] = \Sigma_{z,1} \Sigma_{z,1}^{-1} \underline{u}_i$ cf. page 2
 $= \Sigma_{z,1} \Sigma_{z,1}^{-1} \underline{x}_{z,1}^{(1)}$

$$\begin{aligned} E[\underline{x}_{z,1}^{(2)} | (\underline{x}_{z,1}^{(1)}, U)] &= E[\underline{x}_{z,1}^{(2)} | \underline{x}_{z,1}^{(1)}] - E[w_{21} w_{11}^{-1} \underline{x}_{z,1}^{(1)} | (\underline{x}_{z,1}^{(1)}, U)] \\ &= \underline{\mu}_{z,1}^{(2)} + \Sigma_{z,1} \Sigma_{z,1}^{-1} (\underline{x}_{z,1}^{(1)} - \underline{\mu}_{z,1}^{(1)}) - \Sigma_{z,1} \Sigma_{z,1}^{-1} \underline{x}_{z,1}^{(1)} \\ &= \underline{\mu}_{z,1}^{(2)} - \Sigma_{z,1} \Sigma_{z,1}^{-1} \underline{\mu}_{z,1}^{(1)} = \underline{\mu}_{z,2,1} (= 0, H_0 \text{ true}) \end{aligned}$$

Note that $\text{Var}(\underline{v}_i | U) = \text{Var}(\underline{v}_i | \underline{u}_i) = \Sigma_{z,2,1}$

$$\text{Var}(\underline{x}_{z,1}^{(2)} | (\underline{x}_{z,1}^{(1)}, U)) = \Sigma_{z,2,1} + \sum_i a_i^2 \Sigma_{z,2,1} = (1 + \sum_i a_i^2) \Sigma_{z,2,1}$$

cf. ex. 2.6

$$\underline{x}_{2..1} \mid (\underline{x}^{(1)}, U) \sim N_{d_2}(\underline{\mu}_{2..1}, (1 + \sum_i a_i^2) \Sigma_{22..1})$$

$$\text{Let } \underline{y} = \frac{\underline{x}_{2..1}}{\sqrt{1 + \sum_i a_i^2}}$$

When H_0 is true, $\underline{y} \mid (\underline{x}^{(1)}, U) \sim N_{d_2}(\underline{0}, \Sigma_{22..1})$

The distribution does neither depend on

$(\underline{x}^{(1)}, U)$ nor on $\sum_i a_i^2$, as $\underline{a} = U W_{11}^{-1} \underline{x}^{(1)}$

$$\begin{aligned} \text{Note that } \sum_i a_i^2 &= \underline{a}^T \underline{a} = (U W_{11}^{-1} \underline{x}^{(1)})^T (U W_{11}^{-1} \underline{x}^{(1)}) \\ &= \underline{x}^{(1)T} W_{11}^{-1} U^T U W_{11}^{-1} \underline{x}^{(1)} \\ &= \underline{x}^{(1)T} W_{11}^{-1} \underline{x}^{(1)} = \frac{T_{d_1}^2}{m} \end{aligned}$$

thus \underline{y} is independent of $T_{d_1}^2$

Moreover $W_{22..1}$ is independent of $(T_{d_1}^2, \underline{y})$,

as $(T_{d_1}^2, \underline{y})$ is a function of $(W_{11}, W_{12}, \underline{x})$,

cf. lemma 2.10

$$\left. \begin{array}{l} W_{22..1} \text{ and } (T_{d_1}^2, \underline{y}) \text{ indep.} \\ T_{d_1}^2 \text{ and } \underline{y} \text{ indep.} \end{array} \right\} \Rightarrow W_{22..1}, T_{d_1}^2, \underline{y} \text{ indep.}$$

$$(m-d_1) \underline{y}^T W_{22..1}^{-1} \underline{y} \sim T^2(d_2, m-d_1) \text{ indep. of } T_{d_1}^2$$

$$\text{Now } (m-d_1) \underline{y}^T W_{22..1}^{-1} \underline{y} = \frac{(m-d_1) \underline{x}_{2..1}^T W_{22..1}^{-1} \underline{x}_{2..1}}{1 + \sum_i a_i^2} \frac{m}{m}$$

$$= \frac{(m-d_1)(T_d^2 - T_{d_1}^2)}{m + T_{d_1}^2} \sim \frac{(m-d_1)d_2}{(m-d_1) - d_2 + 1} F(d_2, (m-d_1) - d_2 + 1)$$

independent of $T_{d_1}^2$

$$\Rightarrow \frac{T_d^2 - T_{d_1}^2}{m + T_{d_1}^2} \sim \frac{d_2}{m-d+1} F(d_2, m-d+1)$$

independent of $T_{d_1}^2$

Note the transformation

$$\frac{T_d^2 - T_{d_1}^2}{m + T_{d_1}^2} = \frac{1 + \frac{T_d^2}{m} - \left(1 + \frac{T_{d_1}^2}{m}\right)}{1 + \frac{T_{d_1}^2}{m}} = \frac{1-u}{u}$$

$$\text{where } u = \frac{1 + \frac{T_{d_1}^2}{m}}{1 + \frac{T_d^2}{m}}$$

Skewness and kurtosis for a d -dimensional stochastic vector

$$\text{Skewness: } \beta_{1,d} = E \left[\left((\underline{x} - \underline{\mu})^T \Sigma^{-1} (\underline{y} - \underline{\mu}) \right)^3 \right], \text{ where}$$

$\underline{x}, \underline{y}$ independent, identical distributed

$$\text{Kurtosis: } \beta_{2,d} = E \left[\left((\underline{x} - \underline{\mu})^T \Sigma^{-1} (\underline{x} - \underline{\mu}) \right)^2 \right]$$

Analogue empirical values $(\hat{\Sigma} = \frac{Q}{n})$:

$$\text{Skewness: } \beta_{1,d} = \frac{1}{n^2} \sum_i \sum_j \left((\underline{x}_i - \bar{\underline{x}})^T \hat{\Sigma}^{-1} (\underline{x}_j - \bar{\underline{x}}) \right)^3$$

$$\text{Kurtosis: } \beta_{2,d} = \frac{1}{n} \sum_i \left((\underline{x}_i - \bar{\underline{x}})^T \hat{\Sigma}^{-1} (\underline{x}_i - \bar{\underline{x}}) \right)^2$$