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H. opg. 2.4 24a

$$X = \begin{cases} 0, & \text{når } 0 \leq u \leq \frac{1}{8} \\ 1, & \text{når } \frac{1}{8} < u \leq \frac{1}{2} \\ 2, & \text{når } \frac{1}{2} < u \leq \frac{7}{8} \\ 3, & \text{når } \frac{7}{8} < u \leq 1 \end{cases}, \quad U \sim U[0;1]$$

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H. opg. 2.4 24b

$$X = \begin{cases} 0, & \text{når } 0 \leq u \leq \frac{1}{4} \\ 1, & \text{når } \frac{1}{4} < u \leq \frac{3}{4} \\ 2, & \text{når } \frac{3}{4} < u \leq 1 \end{cases}, \quad U \sim U[0;1]$$

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$Y \sim p(\lambda) = p(1-\lambda)$, dvs. Y kan opfattes som antal begivenheder i $[0; \lambda]$ i en poissonproces m. intensitet 1.

S_n : tid til n te begivenhed

$$S_n = \sum_{k=1}^n X_k, \quad X_k \sim e(1), \quad k=1, \dots, n, \quad \text{uafh.}$$

$$S_n = -\sum_{k=1}^n \ln U_k, \quad U_k \sim U[0;1], \quad k=1, \dots, n, \quad \text{uafh.}$$

$$Y = \min \{n \mid S_{n+1} > \lambda\} = \min \{n \mid S_n > \lambda\} - 1 \\ = \min \{n \mid -\sum_{k=1}^n \ln U_k > \lambda\} - 1$$

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$$N \sim p(\lambda t)$$

$$N = \min \{n \mid -\sum_{k=1}^n \ln U_k > \lambda t\} - 1, \quad \text{H. opg. 3 med } \lambda := \lambda t$$



$$S_j = t U_{(j)}, \quad j=1, \dots, n, \quad U_j \sim U[0;1], \quad j=1, \dots, n, \quad \text{uafh.}$$

H. 'order statistic property'

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$$X \sim nb(r, p)$$

$$X = \sum_{j=1}^r Y_j, \quad Y_j \sim g(p), \quad j=1, \dots, r, \quad \text{uafh.}$$

$$P(Y_j = k) = P(Y_j \leq k) - P(Y_j \leq k-1) = P(Y_j > k-1) - P(Y_j > k)$$

$$Y_j = k, \quad \text{når } (1-p)^k < u < (1-p)^{k-1}, \quad U \sim U[0;1]$$

Bemærk, at $(1-p)^k < u < (1-p)^{k-1} \Leftrightarrow \frac{\ln u}{\ln(1-p)} < k < \frac{\ln u}{\ln(1-p)} + 1$,

dvs. $Y_j = \left\lceil \frac{\ln U}{\ln(1-p)} \right\rceil + 1, \quad U \sim U[0;1].$

$$X = \sum_{j=1}^r \left\lceil \frac{\ln U_j}{\ln(1-p)} \right\rceil + r, \quad U_j \sim U[0;1], \quad j=1, \dots, r, \quad \text{uafh.}$$

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X har tæth. $f(x) = 3x^2$, $0 \leq x \leq 1$

$$F(x) = \int_0^x 3t^2 dt = [t^3]_0^x = x^3, \quad 0 \leq x \leq 1$$

$$u = x^3 \Leftrightarrow x = u^{\frac{1}{3}}$$

$$X = u^{\frac{1}{3}}, \quad u \sim U[0; 1]$$

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$X \sim \text{Cauchy}$, tæth. $f(x) = \frac{1}{\pi(1+x^2)}$, $-\infty < x < \infty$

$$a \quad F(x) = \int_{-\infty}^x \frac{1}{\pi(1+t^2)} dt = \frac{1}{\pi} [\arctan t]_{-\infty}^x = \frac{1}{\pi} \arctan x + \frac{1}{2}$$

$$u = \arctan x + \frac{1}{2} \Leftrightarrow x = \tan\left(\pi\left(u - \frac{1}{2}\right)\right)$$

$$X = \tan\left(\pi\left(u - \frac{1}{2}\right)\right), \quad u \sim U[0; 1]$$

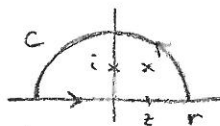
b

$X, Y \sim \text{Cauchy}$ uafh.

$$Z = X + Y$$

$$f_{X+Y}(z) = \int_{-\infty}^{\infty} \frac{1}{\pi(1+x^2)} \frac{1}{\pi(1+(z-x)^2)} dx$$

$$= \frac{1}{\pi^2} \lim_{r \rightarrow \infty} \oint_C \frac{1}{1+w^2} \frac{1}{1+(z-w)^2} dw$$



singulære punkter:

$$w = \pm i \text{ og } w = z \pm i$$

$$= \frac{1}{\pi^2} 2\pi i \left(\frac{1}{2i} \frac{1}{1+(z-i)^2} + \frac{1}{1+(z+i)^2} \frac{1}{2i} \right)$$

$$= \frac{1}{\pi} \left(\frac{1}{z^2 - 2iz} + \frac{1}{z^2 + 2iz} \right)$$

$$= \frac{1}{\pi z} \frac{z + 2i + z - 2i}{z^2 + 4} = \frac{2}{\pi(4 + z^2)}$$

$$f_{\frac{X+Y}{2}}(z) = \frac{2}{\pi(4 + (2z)^2)} |2| = \frac{1}{\pi(1 + z^2)}$$

$$\text{dvs. } \frac{X+Y}{2} \sim \text{Cauchy}$$

Hvis ses, at med $X_1, \dots, X_n \sim \text{Cauchy}$ uafh.,
bliver $\bar{X} \sim \text{Cauchy}$ ($n = 2^m$).

I Cauchy-fordelingen gives det altså ikke
forøget nøjagtighed at tage gennemsnit af
flere observationer. ($\bar{X} \sim \text{Cauchy}$, $n = 2^m$, vises ved
bragtelse af karakteristisk fkt.).

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$$X \sim W(\alpha, \lambda)$$

$$F(x) = 1 - e^{-\lambda x^\alpha}, \quad 0 \leq x < \infty$$

$$u = 1 - e^{-\lambda x^\alpha} \Leftrightarrow x^\alpha = -\frac{1}{\lambda} \ln(1-u) \Leftrightarrow x = \left(-\frac{1}{\lambda} \ln(1-u)\right)^{\frac{1}{\alpha}}$$

$$X = \left(-\frac{1}{\lambda} \ln U\right)^{\frac{1}{\alpha}}, \quad U \sim U[0;1] \quad (U \sim 1-U)$$

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$$X \sim g(r)$$

Fra opg. 2.6.66: $X = [T] + 1, \quad T \sim e(\lambda), \quad r = 1 - e^{-\lambda}$

$$T = -\frac{1}{\lambda} \ln U = \frac{\ln U}{\ln e^{-\lambda}} = \frac{\ln U}{\ln(1-r)}, \quad U \sim U[0;1]$$

$$X = \left\lceil \frac{\ln U}{\ln(1-r)} \right\rceil + 1$$

Samme resultat som mellemresultatet i opg. 5.3.5.

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$$P(X=0) = 0,2$$

$$F_x(x) = 1 - 0,8 e^{-x}, \quad 0 \leq x < \infty$$

$$u = 1 - 0,8 e^{-x} \Leftrightarrow x = -\ln \frac{5(1-u)}{4}$$

$$X = \begin{cases} 0, & \text{når } 0,8 \leq u \leq 1 \\ -\ln \frac{5u}{4}, & \text{når } 0 < u \leq 0,8 \end{cases}, \quad U \sim U[0;1]$$

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$$X \sim e(1), \quad X = -\ln U, \quad U \sim U[0;1]$$

$$\ln X = \ln(-\ln U), \quad U \sim U[0;1]$$

$$E[\ln X] \approx \frac{1}{n} \sum_{k=1}^n \ln(-\ln u_k), \quad u_k \sim U[0;1], \quad k=1, \dots, n, \text{ uafh.}$$

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$$X \sim N(0,1), \text{ tath. } \varphi(x), \quad Y = \sin X$$

$$\int_{-\infty}^{\infty} |\sin x| \varphi(x) dx \leq \int_{-\infty}^{\infty} 1 \varphi(x) dx = 1 \Rightarrow E[|Y|] \text{ eks.}$$

$$E[\sin X] = \int_{-\infty}^{\infty} \sin x \varphi(x) dx = 0$$

simulering ikke nødvendig

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$$X, Y, Z \sim U[0;1] \text{ uafh.},$$

$$P(X+Y+Z \leq 2,5) \approx \frac{1}{n} \sum_{k=1}^n I_k, \quad I_k = \begin{cases} 1, & \text{når } u_1 + u_2 + u_3 \leq 2,5 \\ 0, & \text{når } u_1 + u_2 + u_3 > 2,5 \end{cases}$$

$u_1, u_2, u_3 \sim U[0;1] \text{ uafh.}$

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$$g: [0; 1] \rightarrow \mathbb{R}, \quad I = \int_0^1 g(x) dx$$

$$E[g(U)] = \int_0^1 g(u) \cdot 1 du = I, \quad U \sim U[0; 1]$$

$$I \approx \frac{1}{n} \sum_{k=1}^n g(u_k), \quad u_k \sim U[0; 1], \quad k=1, \dots, n, \quad \text{uafh.}$$

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$$\text{Box-Muller: } \begin{aligned} X &= \sqrt{-2 \ln U} \cos(2\pi V) \\ Y &= \sqrt{-2 \ln U} \sin(2\pi V) \end{aligned}, \quad U, V \sim U[0; 1] \quad \text{uafh.}$$

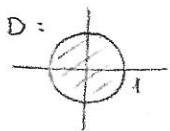
$$\text{Transformation: } X_1 = \sigma_1 X + \mu_1$$

$$\begin{aligned} Y_1 | X_1 &= \sigma_2 \sqrt{1-\rho^2} Y + \mu_{Y|X_1} \\ &= \sigma_2 \sqrt{1-\rho^2} Y + \mu_2 + \rho \frac{\sigma_2}{\sigma_1} (X_1 - \mu_1) \end{aligned}$$

$$(X_1, Y_1) = (\sigma_1 X + \mu_1, \sigma_2 (\rho X + \sqrt{1-\rho^2} Y) + \mu_2)$$

$$(X_1, Y_1) \sim N_2(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$$

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$$f(r, \theta) = \frac{1}{\pi}, \quad \begin{aligned} 0 \leq r \leq 1 \\ 0 \leq \theta < 2\pi \end{aligned}, \quad \begin{aligned} x &= r \cos \theta \\ y &= r \sin \theta \end{aligned}$$

$$f_R(r) = \int_0^{2\pi} \frac{1}{\pi} r d\theta = \frac{r}{\pi} [0]_0^{2\pi} = 2r$$

$$F_R(r) = \int_0^r 2t dt = 2 \left[\frac{t^2}{2} \right]_0^r = r^2, \quad u = r^2 \Rightarrow r' = \sqrt{u}$$

$$f_\Theta(\theta) = \int_0^1 \frac{1}{\pi} r dr = \frac{1}{\pi} \left[\frac{r^2}{2} \right]_0^1 = \frac{1}{2\pi} \Rightarrow \Theta \sim U[0; 2\pi]$$

$$(X, Y) = (\sqrt{U} \cos(2\pi V), \sqrt{U} \sin(2\pi V))$$

$$U, V \sim U[0; 1] \quad \text{uafh.}$$

$$(X, Y) \sim U[D]$$

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$$X, \text{ tæth. } f(x) = \frac{4}{\sqrt{\pi}} x^2 e^{-x^2}, \quad 0 \leq x < \infty$$

$$Y \sim \Gamma(3, 1), \quad g(y) = \frac{1}{2} y^2 e^{-y}$$

$$\frac{f(y)}{c g(y)} = \frac{\frac{4}{\sqrt{\pi}} y^2 e^{-y^2}}{c \frac{1}{2} y^2 e^{-y}} = \frac{1}{c} \frac{8}{\sqrt{\pi}} e^{-(y^2 - y + \frac{1}{4})} e^{\frac{1}{4}}$$

$$= \frac{1}{c} \frac{8e^{\frac{1}{4}}}{\sqrt{\pi}} e^{-(y - \frac{1}{2})^2} \leq 1 \quad \text{for } c = \frac{8e^{\frac{1}{4}}}{\sqrt{\pi}} \approx 5.80$$

fortsætter

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fortsatt

$$Y = \sum_{k=1}^3 Z_k, \quad Z_k \sim e(1), \quad Z_k = -\ln V_k, \quad k=1, 2, 3$$

uafh.

$$Y = - \sum_{k=1}^3 \ln V_k$$

$$U \sim U[0, 1]$$

$$X = y, \text{ n\u00e5r } u < e^{-(y-\frac{1}{2})^2},$$

ellers afvis