Prediction

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April 3, 2012

Decomposition of *X*:

$$X = \mathbb{E}[X|Y] + (X - \mathbb{E}[X|Y])$$

where predictor $\mathbb{E}[X|Y]$ and prediction error $X - \mathbb{E}[X|Y]$ uncorrelated. Hence

$$\mathbb{V}$$
ar $X = \mathbb{V}$ ar $\mathbb{E}[X|Y] + \mathbb{V}$ ar $(X - \mathbb{E}[X|Y])$

Note:

$$\mathbb{E} \mathbb{V}\mathrm{ar}(X|Y) = \mathbb{V}\mathrm{ar}(X - \mathbb{E}[X|Y])$$

Optimal prediction

Consider two random variables X and Y. Let $\tilde{X} = f(Y)$ be a predictor of X based on Y. The mean square prediction error is

$$\mathbb{E}(X-\tilde{X})^2$$

The predictor that minimizes the prediction error is the conditional expectation $\mathbb{E}[X|Y]$:

$$\begin{split} \mathbb{E}(X-\tilde{X})^2 &= \\ \mathbb{E}(X-\mathbb{E}[X|Y])^2 + \mathbb{E}(\mathbb{E}[X|Y]-\tilde{X})^2 + 2\mathbb{C}\mathrm{ov}(X-\mathbb{E}[X|Y],\mathbb{E}[X|Y]-\tilde{X}) \end{split}$$
 and

$$\begin{split} \mathbb{C}\text{ov}(X - \mathbb{E}[X|Y], \mathbb{E}[X|Y] - \tilde{X}) &= \\ \mathbb{C}\text{ov}(\mathbb{E}[X - \mathbb{E}[X|Y]|Y], \mathbb{E}[\mathbb{E}[X|Y] - \tilde{X}|Y]) + \\ \mathbb{E}\mathbb{C}\text{ov}(X - \mathbb{E}[X|Y], \mathbb{E}[X|Y] - \tilde{X}|Y) &= 0 \end{split}$$

("Pythagoras for random variables", E[X|Y] "projection" of X on Y)

Pythagoras and conditional expectation

Space of real random variables with finite variance may be viewed as a vector space with inner product and (L_2) norm

$$\langle X, Y \rangle = \mathbb{C}\text{ov}[X, Y] \quad ||X|| = \sqrt{\mathbb{V}\text{ar}X}$$

Two random variables are orthogonal if they are uncorrelated. Pythagoras:

$$||X||^2 = ||\mathbb{E}[X|Y]||^2 + ||X - \mathbb{E}[X|Y]||^2$$

 $\mathbb{E}[X|Y]$ may be viewed as projection of X on Y.

Conditional expectation in multivariate normal distribution

Consider jointly normal random vectors X and Y with mean vector

$$\mu = (\mu_X, \mu_Y)$$

and covariance matrix

$$\Sigma = egin{bmatrix} \Sigma_X & \Sigma_{XY} \ \Sigma_{YX} & \Sigma_Y \end{bmatrix}$$

Then (provided Σ_Y invertible)

$$X|Y = y \sim N(\mu_X + \Sigma_{XY}\Sigma_Y^{-1}(y - \mu_Y), \Sigma_X - \Sigma_{XY}\Sigma_Y^{-1}\Sigma_{YX})$$

Proof:

$$X = Z + R$$

where $Z = \mu_X + \sum_{XY} \sum_{YY}^{-1} (Y - \mu_Y)$ and R = X - Z are independent and R is independent of Y.

Prediction in linear mixed model (Jiang page 75)

Let $\alpha \sim N(0, G)$ and $Y | \alpha \sim N(X\beta + Z\alpha, R)$.

Then $\mathbb{C}\text{ov}[\alpha, Y] = GZ^{\mathsf{T}}$ and $\mathbb{V}\text{ar} Y = ZGZ^{\mathsf{T}} + R$.

$$\mathbb{E}[\alpha|Y] = GZ^{\mathsf{T}}V^{-1}(Y - X\beta)$$

NB:

$$GZ^{\mathsf{T}}(ZGZ^{\mathsf{T}}+R)^{-1}=(G^{-1}+Z^{\mathsf{T}}R^{-1}Z)^{-1}Z^{\mathsf{T}}R^{-1}$$

e.g. useful if G^{-1} is sparse (like AR-model).

Conditional simulation using prediction

Suppose we wish to simulate X|Y=y. By previous decomposition

$$X = Z + R$$

we can simulate (X^*, Y^*) and compute Z^* and $R^* = X^* - Z^*$.

Then our conditional simulation is

$$\mathbb{E}[X|Y=y]+R^*$$

Advantageous if it is easier to simulate (X^*, Y^*) and compute Z^* than simulate directly from conditional distribution of X|Y=y

(e.g. if Σ_X and Σ_Y easier than $\Sigma_X - \Sigma_{XY} \Sigma_Y^{-1} \Sigma_{YX}$)

IQ example (Jiang)

$$Y = \mu + \alpha + \epsilon$$

where standard deviation of α and ϵ are 15 and 5 and $\mu=$ 100.

Given
$$Y = 130$$
, $\mathbb{E}[\mu + \alpha | Y = 130] = 127$.

Example of shrinkage to the mean.

BLUP

Consider random vectors X and Y with with mean vector

$$\mu = (\mu_X, \mu_Y)$$

and covariance matrix

$$\Sigma = egin{bmatrix} \Sigma_X & \Sigma_{XY} \ \Sigma_{YX} & \Sigma_Y \end{bmatrix}$$

Then best linear unbiased predictor (BLUP) of X given Y is

$$\hat{X} = \mu_X + \Sigma_{XY} \Sigma_Y^{-1} (Y - \mu_Y)$$

 $(\mathbb{E}\hat{X} = \mathbb{E}X$, linear in Y and smallest prediction error among linear unbiased predictors)

Covariance matrix of BLUP:

$$\mathbb{C}\mathrm{ov}\hat{X} = \Sigma_{XY}\Sigma_{Y}^{-1}\Sigma_{YX}$$

EBLUP and **EBLUE**

Typically covariance matrix depends on unknown parameters.

EBLUPS are obtained by replacing unknown variance parameters by their estimates (similar for EBLUE).

Prediction of mixed effect - Henderson's equations

Given β , BLUP of

$$b^{\mathsf{T}}\beta + a^{\mathsf{T}}\alpha$$

is

$$b^{\mathsf{T}}\beta + a^{\mathsf{T}}GZ^{\mathsf{T}}V^{-1}(Y - X\beta)$$

Typically β is unknown. Then BLUP is

$$b^{\mathsf{T}}\hat{\beta} + a^{\mathsf{T}}GZ^{\mathsf{T}}V^{-1}(Y - X\hat{\beta})$$

where $\hat{\beta}$ is BLUE (Harville, 1991)

NB: $(\hat{\alpha}, \hat{\beta})$ solution of Henderson's mixed model equations.

Model assessment

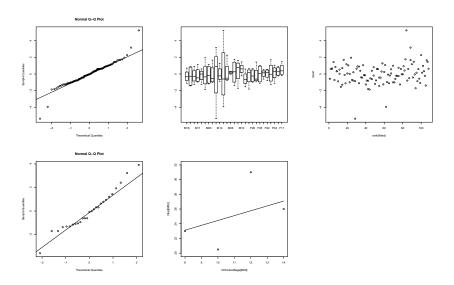
Make histograms, qq-plots etc. for EBLUPs of ϵ and α .

May be advantageous to consider standardized EBLUPS. Standardized BLUP is

$$[\mathbb{C}\mathrm{ov}\hat{\alpha}]^{-1/2}\hat{\alpha}$$

Example: prediction of random intercepts and slopes in orthodont data

```
ort7=lmer(distance~age+factor(Sex)+(1|Subject),data=Orthode
#check of model ort7
#residuals
res=residuals(ort7)
qqnorm(res)
qqline(res)
#outliers occur for subjects MO9 and M13
#plot residuals against subjects
boxplot(resort~Orthodont$Subject)
#plot residuals against fitted values
fitted=fitted(ort7)
plot(rank(fitted),resort)
```



```
#extract predictions of random intercepts
raneffects=ranef(ort7)
#qqplot of random intercepts
qqnorm(ranint[[1]])
qqline(ranint[[1]])
#plot for subject M09
M09=Orthodont$Subject=="M09"
plot(Orthodont$subject=="M09],fitted[M09],type="1",ylim=c(20,32)]
points(Orthodont$age[M09],Orthodont$distance[M09])
```