

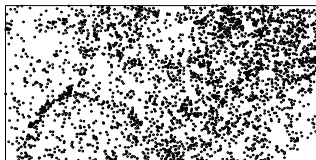
Decomposition of variance for spatial Cox processes

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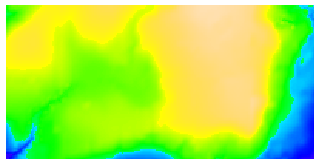
Joint work with Abdollah Jalilian and Yongtao Guan

March 15, 2011

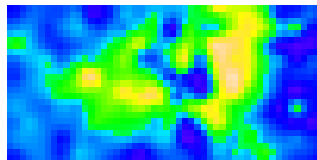
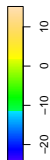
Tropical rain forest example: *Capparis Frondosa*



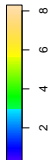
- ▶ observation window W
= 1000 m \times 500 m
- ▶ seed dispersal \Rightarrow *clustering*
- ▶ environment \Rightarrow *inhomogeneity*



Elevation



Potassium content in soil.



How much variation due to environmental variables and how much due to seed dispersal ?

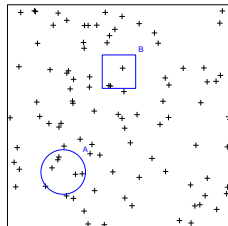
Framework: spatial Cox point processes.

Poisson and Cox processes

Poisson process with intensity function $\rho(\cdot)$:

counts $N(B)$ independent
and Poisson distributed with mean

$$\mathbb{E}N(B) = \int_B \rho(u) du$$



Cox process \mathbf{X} with random intensity function Λ : conditional on $\Lambda = \lambda$, \mathbf{X} Poisson process with intensity function λ .

Additive model for Λ

Additive model:

$$\Lambda(u) = \tilde{Z}(u) + \Lambda_0(u)$$

\tilde{Z} contribution from environment:

$$\tilde{Z}(u) = \beta Z(u)^T \quad Z(u) = (Z_1(u), \dots, Z_p(u))$$

Λ_0 : structured variation (e.g. seed dispersal) not due to environment.

Assume \tilde{Z} and Λ_0 independent non-negative and stationary.

Cox process \mathbf{X} superposition of independent Cox processes with random intensity functions $\tilde{Z}(u)$ and $\Lambda_0(u)$.

Log-linear model

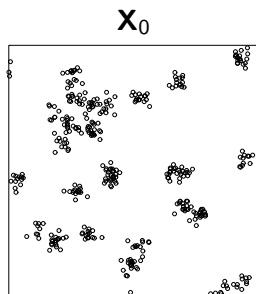
$$\Lambda(u) = \Lambda_0(u) \exp[\tilde{Z}(u)]$$

Interpretation in terms of survival of seedlings:

\mathbf{X}_0 seedlings: stationary Cox process with random intensity function Λ_0 .

\mathbf{X} thinning of \mathbf{X}_0 with survival depending on environment \tilde{Z} .

Example: Cox/cluster process: Inhomogeneous Thomas process

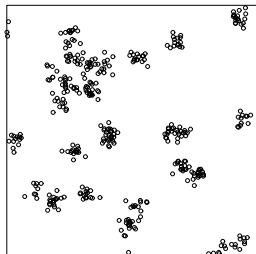


Λ_0 shot-noise process $\Rightarrow \mathbf{X}_0$ cluster process:

Offspring distributed around Poisson parents according to Gaussian density

Example: Cox/cluster process: Inhomogeneous Thomas process

\mathbf{X}_0



Λ_0 shot-noise process $\Rightarrow \mathbf{X}_0$ cluster process:

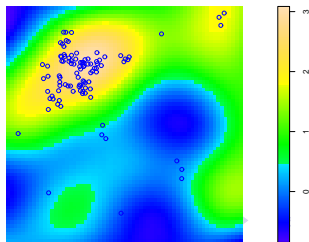
Offspring distributed around Poisson parents according to Gaussian density

Inhomogeneity: offspring in \mathbf{X}_0 survive according to probability

$$p(u) \propto \exp[\beta Z(u)^T]$$

depending on covariates (independent thinning).

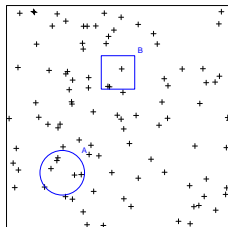
\mathbf{X}



Decomposition of variance for a count

Prediction of count $N(B)$ given Λ :

$$\hat{N}(B) = \mathbb{E}[N(B)|\Lambda] = \int_B \Lambda(u) du$$



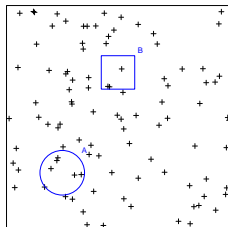
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variation =



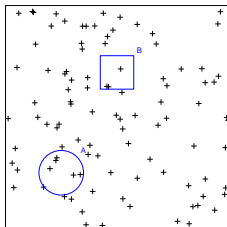
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Decomposition of variance for a count

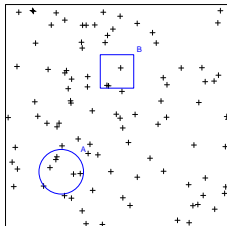
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$$\hat{N}(B) = \mathbb{E}[N(B)|\Lambda] = \int_B \Lambda(u) du$$

$$\text{Var}N(B) = \text{Var}\hat{N}(B) + \text{Var}[N(B) - \hat{N}(B)]$$

variation = variation Λ + 'Poisson noise'

$$\text{Further, } \hat{\Lambda}(u) = \mathbb{E}[\Lambda(u)|Z]$$



Decomposition of variance for a count

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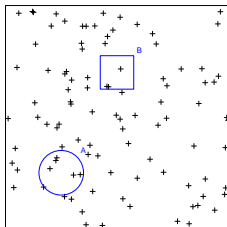
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structured variation =

$$\text{SST} =$$



Decomposition of variance for a count

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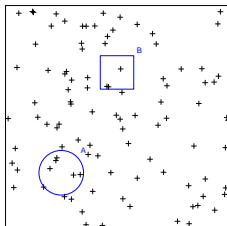
variation = variation Λ + 'Poisson noise'

Further, $\hat{\Lambda}(u) = \mathbb{E}[\Lambda(u)|Z]$

$$\text{Var}\Lambda(u) = \text{Var}\hat{\Lambda}(u)$$

structured variation = variation due to environment

$$\text{SST} = \text{SSR}$$



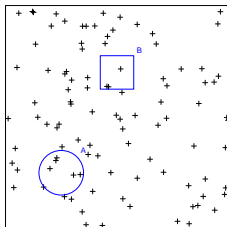
Decomposition of variance for a count

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Further, $\hat{\Lambda}(u) = \mathbb{E}[\Lambda(u)|Z]$

$$\text{Var}\Lambda(u) = \text{Var}\hat{\Lambda}(u) + \text{Var}[\Lambda(u) - \hat{\Lambda}(u)]$$

structured variation = variation due to environment + other sources

$$\text{SST} = \text{SSR} + \text{SSE}$$

Measure of influence of environmental covariates Z :

$$R^2 = \frac{SSR}{SST} = \frac{\text{VarE}[\Lambda(u)|Z]}{\text{Var}\Lambda(u)}$$

(right hand side does not depend on u in case of stationary environment)

R^2 for additive and log-linear models

Additive:

$$R^2 = \frac{\sigma_{\tilde{Z}}^2}{\sigma_{\tilde{Z}}^2 + \sigma_0^2}$$

$$\sigma_{\tilde{Z}}^2 = \text{Var} \tilde{Z}(u) \quad \text{and} \quad \sigma_0^2 = \text{Var} \Lambda_0(u)$$

Log-linear:

$$R^2 = \frac{\sigma_{\exp \tilde{Z}}^2}{\sigma_{\exp \tilde{Z}}^2 + \sigma_0^2 [\sigma_{\exp \tilde{Z}}^2 + \mu_{\exp \tilde{Z}}^2]}$$

$$\sigma_{\exp \tilde{Z}}^2 = \text{Var} \exp[\tilde{Z}(u)] \quad \text{and} \quad \mu_{\exp \tilde{Z}} = \mathbb{E} \exp[\tilde{Z}(u)]$$

Estimation: environmental variances

Z observed on grid $G = \{u_i\}_{i=1,\dots,M}$

Simple empirical estimates, e.g.

$$\hat{\sigma}_{\tilde{Z}}^2 = \frac{1}{M} \left\{ \sum_{u \in G} \hat{\tilde{Z}}(u)^2 - \frac{[\sum_{u \in G} \hat{\tilde{Z}}(u)]^2}{M} \right\}$$

where $\hat{\tilde{Z}}(u) = \hat{\beta}Z(u)^T$.

Estimation: β

Estimate β conditioning on Z .

First-order log composite likelihood:

$$\text{CL}_1(\beta) = \sum_{u \in \mathbf{X}} \log \rho(u|Z; \beta) - \int_W \rho(u|Z; \beta) du$$

$\rho(\cdot|Z, \beta) = \mathbb{E}[\Lambda(u)|Z]$ intensity function for $\mathbf{X}|Z$.

Parametric models for $c_0(u - v) = \text{Cov}[\Lambda_0(u), \Lambda_0(v)]$

Any positive definite function is a covariance function but not necessarily for a non-negative random process Λ_0 . Use covariance functions from explicit constructions of Λ_0 .

Log-Gaussian:

$$\Lambda_0(u) = \exp[Y(u)]$$

where Y Gaussian field.

Covariance:

$$c_0(u - v) = \rho_0^2 \{ \exp[\text{Cov}(Y(u), Y(v))] - 1 \}$$

where $\rho_0 = \mathbb{E}\Lambda_0(u)$.

Shot-noise:

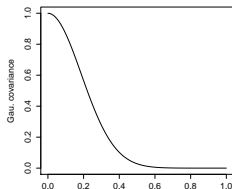
$$\Lambda_0(u) = \sum_{v \in C} \alpha k(u - v)$$

where C homogeneous Poisson with intensity κ and $k(\cdot)$ probability density.

$$c_0(u - v) = \kappa \alpha^2 \int_{\mathbb{R}^2} k(w) k(w + v - u) dw$$

Example: kernel k Gaussian density \Rightarrow modified Thomas/'Gaussian' covariance function:

$$c_0(h) = \sigma_0^2 \exp[-(\|h\|/\eta)^2]$$



Bessel shot-noise/Matérn covariance

Λ_0 Bessel shot-noise process: sum of $K_{(\nu-1)/2}$ Bessel densities centered around points of homogeneous Poisson process.

Matérn covariance function:

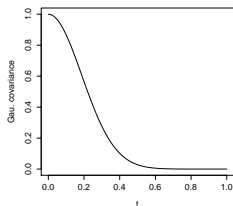
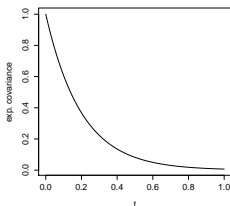
$$c_0(h) = \sigma_0^2 \frac{(\|h\|/\eta)^\nu K_\nu(\|h\|/\eta)}{2^{\nu-1} \Gamma(\nu)}$$

$\nu = 1/2$: exponential model

' $\nu = \infty$ ': 'Gaussian'

$$c_0(h) = \sigma_0^2 \exp(-\|h\|/\eta)$$

$$c_0(h) = \sigma_0^2 \exp[-(\|h\|/\eta)^2]$$



Estimation of ψ ($c_0(\cdot) = c_0(\cdot; \psi)$)

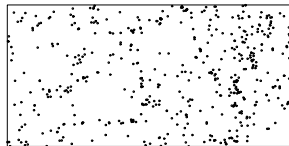
Second-order log composite likelihood (given $\hat{\beta}$, conditioning on Z):

$$\begin{aligned} \text{CL}_2(\psi|\hat{\beta}) &= \sum_{\substack{u, v \in \mathbf{X} \\ \|u-v\| \leq R}}^{\neq} \log \rho^{(2)}(u, v|Z; \hat{\beta}, \psi) \\ &\quad - \iint_{\|u-v\| \leq R} \rho^{(2)}(u, v|Z; \hat{\beta}, \psi) du dv \end{aligned}$$

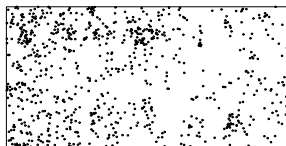
$\rho^{(2)}(u, v|Z; \beta, \psi) = \mathbb{E}[\Lambda(u)\Lambda(v)|Z]$ second-order product density for $\mathbf{X}|Z$.

Three species with different modes of seed dispersal:

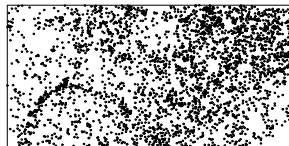
Acalypha Diversifolia explosive capsules



Loncocharpus Heptaphyllus wind



Capparis Frondosa bird/mammal

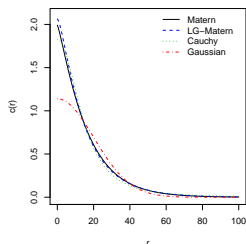


Results for rain forest data

Species	model for Λ	c_0	$\Delta\text{CL}_2(\hat{\psi} \hat{\beta})$	R^2
Acalypha	log-linear	'Gaussian'	402	0.01
		Matérn	421	0.02
	additive	'Gaussian'	0	0.01
		Matérn	18	0.01
Loncocharpus	log-linear	'Gaussian'	877	0.17
		Matérn	925	0.10
	additive	'Gaussian'	0	0.11
		Matérn	26	0.07
Capparis	log-linear	'Gaussian'	4482	0.37
		Matérn	4814	0.19
	additive	'Gaussian'	0	0.23
		Matérn	454	0.16

Some conclusions

Covariance
functions
for *loncocharpus*



Best fit with Matérn (heavy tails for covariance/cluster density).

Best fit with log-linear model
(interpretation in terms of survival).

Largest R^2 for *Capparis* (bird/mammal seed dispersal), smallest for *Acalypha* (explosive capsules).