

# Computation of the likelihood function for GLMMs

Rasmus Waagepetersen  
Department of Mathematics  
Aalborg University  
Denmark

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- ▶ likelihood for GLMM
- ▶ penalized quasi-likelihood estimation
- ▶ Laplace approximation
- ▶ Gaussian quadrature
- ▶ case study of non-linear mixed effects model

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## Generalized linear mixed effects models

Consider stochastic variable  $\mathbf{Y} = (Y_1, \dots, Y_n)$  and random effects  $\mathbf{U}$ .

Two step formulation of GLMM:

- ▶  $\mathbf{U} \sim N(0, \Sigma)$ .
- ▶ Given realization  $\mathbf{u}$  of  $\mathbf{U}$ ,  $Y_i$  independent and each follows density  $f_i(\mathbf{y}|\mathbf{u})$  with mean  $\mu_i = g^{-1}(\eta_i)$  and linear predictor  $\eta = X\beta + Z\mathbf{u}$ .

I.e. conditional on  $\mathbf{U}$ ,  $Y_i$  follows a generalized linear models.

NB: GLMM specified in terms of marginal density of  $\mathbf{U}$  and conditional density of  $\mathbf{Y}$  given  $\mathbf{U}$ . But the likelihood is the marginal density of  $f(\mathbf{y})$  which can be hard to evaluate !

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## Likelihood for generalized linear mixed model

For normal linear mixed models we could compute the marginal distribution of  $\mathbf{Y}$  directly as a multivariate normal. That is,  $f(\mathbf{y})$  is a density of a multivariate normal distribution.

For a generalized linear mixed model it is difficult to evaluate the integral:

$$f(\mathbf{y}) = \int_{\mathbb{R}^m} f(\mathbf{y}, \mathbf{u}) d\mathbf{u} = \int_{\mathbb{R}^m} f(\mathbf{y}|\mathbf{u}) f(\mathbf{u}) d\mathbf{u}$$

since  $f(\mathbf{y}|\mathbf{u})f(\mathbf{u})$  is a very complex function.

Today: numerical methods for evaluating likelihood of GLMM.

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## Non-normal example: logistic regression with random intercepts

$$\begin{aligned}
 U_j &\sim N(0, \tau^2), \quad j = 1, \dots, m \\
 Y_j | U_j = u_j &\sim \text{binomial}(n_j, p_j) \\
 \log(p_j / (1 - p_j)) &= \eta_j = \beta + U_j \\
 p_j &= \exp(\eta_j) / (1 + \exp(\eta_j))
 \end{aligned}$$

Conditional density:

$$f(y|u; \beta) = \prod_j p_j^{y_j} (1 - p_j)^{1 - y_j} = \prod_j \frac{\exp(\beta + u_j)^{y_j}}{(1 + \exp(\beta + u_j))^{n_j}}$$

Likelihood function ( $u = (u_1, \dots, u_m)$ )

$$\int_{\mathbb{R}^m} f(y|u; \beta) f(u; \tau^2) du = \prod_j \int_{\mathbb{R}} \frac{\exp(\beta + u_j)^{y_j}}{(1 + \exp(\beta + u_j))^{n_j}} \frac{\exp(-u_j^2 / (2\tau^2))}{\sqrt{2\pi\tau^2}} du_j$$

Integrals can not be evaluated in closed form.

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## One-dimensional case

Compute

$$L(\theta) = \int_{\mathbb{R}} f(y|u; \beta) f(u; \tau^2) du$$

Possibilities:

- ▶ Laplace approximation.
- ▶ Numerical integration/quadrature (e.g. Gaussian quadrature as in PROC NL MIXED (SAS) or GLLAM (STATA)) (one level of random effects, dimensions one or two).

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## Laplace approximation

Let  $g(u) = \log(f(y|u)f(u))$  and choose  $\hat{u}$  so  $g'(\hat{u}) = 0$  ( $\hat{u} = \arg \max g(u)$ ).

Taylor expansion around  $\hat{u}$ :

$$\begin{aligned}
 g(u) &\approx \tilde{g}(u) = \\
 g(\hat{u}) &+ (u - \hat{u})g'(\hat{u}) + \frac{1}{2}(u - \hat{u})^2 g''(\hat{u}) = g(\hat{u}) - \frac{1}{2}(u - \hat{u})^2 (-g''(\hat{u}))
 \end{aligned}$$

i.e.  $\exp(\tilde{g}(u))$  proportional to normal density  $N(\mu_{LP}, \sigma_{LP}^2)$ ,  
 $\mu_{LP} = \hat{u}$   $\sigma_{LP}^2 = -1/g''(\hat{u})$ .

$$\begin{aligned}
 L(\theta) &= \int_{\mathbb{R}} \exp(g(u)) du \approx \int_{\mathbb{R}} \exp(\tilde{g}(u)) du \\
 &= \exp(g(\hat{u})) \int_{\mathbb{R}} \exp\left(-\frac{1}{2\sigma_{LP}^2}(u - \mu_{LP})^2\right) du = \exp(g(\hat{u})) \sqrt{2\pi\sigma_{LP}^2}
 \end{aligned}$$

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Laplace approximation also possible for higher dimensions (multivariate Taylor expansion).

NB:

$$f(u|y) = f(y|u)f(u)/f(y) \propto \exp(g(u)) \approx \text{const} \exp\left(-\frac{1}{2\sigma_{LP}^2}(u - \mu_{LP})^2\right)$$

where  $\mu_{LP} = \hat{u}$   $\sigma_{LP}^2 = -1/g''(\hat{u})$ .

Hence

$$U|Y = y \approx N(\mu_{LP}, \sigma_{LP}^2)$$

Note:  $\mu_{LP}$  is mode of conditional distribution - used for prediction of random effects in `lmer (ranef())`.

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## Adaptive Gaussian Quadrature

Gauss-Hermite quadrature (numerical integration) is

$$\int f(x)\phi(x)dx \approx \sum_{i=1}^n w_i f(x_i)$$

where  $\phi$  is the standard normal density and  $(x_i, w_i), i = 1, n$  are certain arguments and weights which can be looked up in a table.

We can replace  $\approx$  with  $=$  whenever  $f$  is a polynomial of degree  $2n - 1$  or less.

Adaptive Gauss-Hermite quadrature:

$$\int f(y|u)f(u)du \approx \int \frac{f(y|u)f(u)}{\phi(u; \mu_{LP}, \sigma_{LP}^2)} \phi(u; \mu_{LP}, \sigma_{LP}^2) du = \int \frac{f(y|\sigma_{LP}x + \mu_{LP})f(\sigma_{LP}x + \mu_{LP})}{\phi(x)} \sigma_{LP} \phi(x) dx$$

(change of variable:  $x = (u - \mu_{LP})/\sigma_{LP}$ )

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Advantage

$$\frac{f(y|u)f(u)}{\phi(u; \mu_{LP}, \sigma_{LP}^2)} = \frac{f(y|\sigma_{LP}x + \mu_{LP})f(\sigma_{LP}x + \mu_{LP})}{\phi(x)} \quad x = (u - \mu_{LP})/\sigma_{LP}$$

close to constant ( $f(y)$ ) – hence G-H quadrature very accurate.

GH scheme with  $n = 5$ :

x		2.020	0.959	0.0000000	-0.959	-2.020	(computed using ghq in library g1mmML).
w		0.011	0.222	0.533	0.222	0.011	

(GH schemes for  $n = 5$  and  $n = 10$  available on web page)

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## Prediction of random effects for GLMM

Conditional mean

$$\mathbb{E}[U|Y = y] = \int uf(u|y)du$$

is minimum mean square error predictor, i.e.

$$\mathbb{E}(U - \hat{U})^2$$

is minimal with  $\hat{U} = H(Y)$  where  $H(y) = \mathbb{E}[U|Y = y]$

Difficult to analytically evaluate

$$\mathbb{E}[U|Y = y] = \int uf(y|u)f(u)/f(y)du$$

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## Computation of conditional expectations (prediction)

$$\mathbb{E}[U|Y = y] = \int u \frac{f(y|u)f(u)}{f(y)} du = \frac{1}{f(y)} \int (\sigma_{LP}x + \mu_{LP}) \frac{f(y|\sigma_{LP}x + \mu_{LP})f(\sigma_{LP}x + \mu_{LP})}{\phi(x)} \sigma_{LP} \phi(x) dx$$

Note:

$$(\sigma_{LP}x + \mu_{LP}) \frac{f(y|\sigma_{LP}x + \mu_{LP})f(\sigma_{LP}x + \mu_{LP})}{\phi(x)} \sigma_{LP}$$

behaves like a first order polynomial in  $x$  - hence GH still accurate.

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## Penalized quasi-likelihood

One solution: do not use likelihood function but something simpler.

$$\theta = (\beta, \tau^2)$$

PQL estimates  $\hat{\theta}$  and  $\hat{u}$  maximize joint density

$$f(y, u; \theta) = f(y|u; \beta)f(u; \tau^2).$$

PQL estimates less accurate than ML.

Asymptotic results require increasing number of observations for each random effect.

Implemented in `lmer` and SAS macro `glimmix`.

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## Difficult cases

- ▶ correlated random effects
- ▶ crossed random effects
- ▶ nested random effects

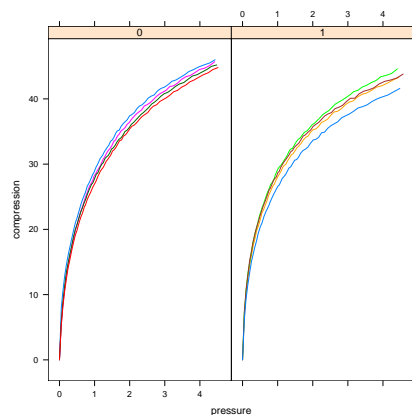
Not possible to factorize likelihood into low-dimensional integrals – hence numerical quadrature not applicable.

PQL and Laplace-approximation still applicable (`lmer()`).

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## Case study: non-linear mixed effects model for cow mats data

Compression vs. pressure for two brands of mats



Non-linear relation

$$y = \text{mmf}(x) = \frac{ab + cx^d}{b + x^d},$$

Random variation between mats of same brand, small measurement noise.

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Estimation of non-linear model with fixed effects:

```
nlsfit=nls(nedtryk~mmf(tryk,a,b,c,d),start=c(a=0.1,b=1.670,c=80,d=0.6),data=mattressdata1)
```

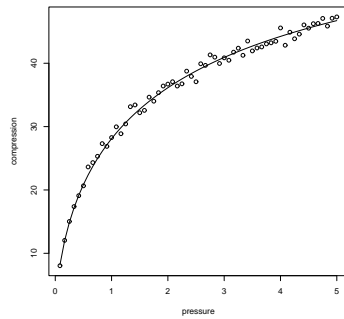
Estimation of non-linear model with  $a$ ,  $b$ ,  $c$  as random effects:

```
nlmerfit=nlmer(nedtryk~mmfnlmer(tryk,a,b,c,d)~(a|matno)+(b|matno)+(c|matno),mattressdata1,start=c(a=0.04,b=1.64,c=74,d=0.64))
```

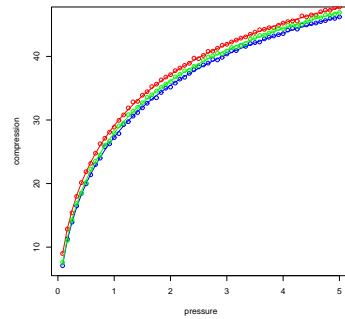
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Simulated data from the two models:

Fixed effects: residual standard error 0.72



With random effects: residual standard error 0.17



Std. err. for  $a, b, c$  are 0.64, 0.05 and 0.14

Random effects model gives much better representation of variability in data.

NB: to assess influence of variability of different parameters we need to look at partial derivatives (sensitivities) wrt. these parameters.

## Exercises

See exercises sheet on webpage.